

# Time Series Analysis (& SI)—191571090

(The first 1.5 hours it is CLOSED BOOK. The second 1.5 hours is OPEN BOOK.)

Date: 4-nov-2016  
Place: Sports center  
Time: 08:45–11:45

1. **CLOSED BOOK:** Three questions:

- (a) Let  $r(0) = 3, r(\pm 1) = 2, r(\pm 2) = 1$  and all other  $r(k) = 0$ . Is this a covariance function?
- (b) When do we say that an estimator is “efficient”?
- (c) Is it correct to claim that  $U_t$  defined as  $U_t = B(q)\epsilon_t$  is in general not a good choice as input for system identification if  $B(z)$  has a zero on the unit circle (i.e.  $B(z_1) = 0$  for some  $z_1 \in \mathbb{C}$  with  $|z_1| = 1$ )?

2. **CLOSED BOOK:** Consider the ARMA process

$$X_t = aX_{t-2} + \epsilon_t + b\epsilon_{t-1}.$$

(Notice the “ $t-2$ ”.)

- (a) For which  $a, b$  is  $X_t$  asymptotically wide sense stationary?
  - (b) For which  $a, b$  is the scheme invertible.
  - (c) Assume it is asymptotically WSS and invertible. Determine the 1-step ahead predictor
  - (d) Assume it is asymptotically WSS and invertible. Determine the 2-step ahead predictor
  - (e) Actually the standard formula for the 1-step ahead predictor assumes that  $\epsilon_t$  is zero mean. How would you modify the above 1-step ahead predictor if  $\mu := \mathbb{E}\epsilon_t$  is nonzero?
  - (f) Determine all  $a, b$  for which  $X_t$  is white noise.
3. **CLOSED BOOK:** The Gauss-Markov theorem says that the best linear unbiased estimator (BLUE)  $\hat{\theta} = KX$  of  $\theta$  where

$$X = W\theta + \epsilon,$$

equals the least squares solution if  $\epsilon$  is zero mean and has covariance matrix  $\sigma^2 I$ . What is the BLUE  $\hat{\theta} = KX$  in case  $\epsilon$  is zero mean but has a diagonal covariance matrix of the form

$$R_\epsilon = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 2 & 0 & 0 \\ 0 & 0 & 3 & 0 \\ 0 & 0 & 0 & \ddots \end{bmatrix}.$$

4. **CLOSED BOOK:** Given data  $x_0, \dots, x_{N-1}$  is it numerically easier to estimate the parameters of an AR-model or of an MA-model?

5. **OPEN BOOK:** In § 6.1 it is proved that the variance of the sample mean  $\hat{m}_N = \frac{1}{N}(x_0 + \dots + x_{N-1})$  for wide-sense stationary processes  $X_t$  is

$$\frac{1}{N} \sum_{k=-N+1}^{N-1} (1 - |k|/N) r(k).$$

The problem, of course, is that we do not know  $r(k)$ . Some researchers suggest to estimate this variance as

$$\frac{1}{N} \sum_{k=-M}^M (1 - |k|/N) \hat{r}_N(k)$$

for some suitable choice of  $M$ . How would you choose  $M$ ? (Explain your choice of  $M$ .)

6. **OPEN BOOK:** Suppose  $X_t$  is zero mean white noise. The periodogram  $p_N(\omega)$  is defined as

$$P_N(\omega) = \frac{1}{N} \left| \sum_{t=0}^{N-1} X_t e^{-i\omega t} \right|^2.$$

For each  $\omega$  compute  $E(P_N(\omega))$ .

7. **OPEN BOOK:** In the lecture notes the system  $h_t$  is resolved from Equation (8.10) through Fourier transformation, but it can also be done in time domain if we assume that the system has “finite impulse response” meaning that

$$y_t = \sum_{m=0}^{m=M} h_m u_{t-m} + v_t$$

for some finite  $M > 0$ . That is,  $h_t = 0$  for all  $t > M$ .

- (a) Show that the Equation (8.10) for finite impulse response systems implies that

$$\begin{bmatrix} r_u(0) & r_u(1) & \dots & r_u(M) \\ r_u(1) & r_u(0) & \dots & r_u(M-1) \\ \vdots & \vdots & \ddots & \vdots \\ r_u(M) & r_u(M-1) & \dots & r_u(0) \end{bmatrix} \begin{bmatrix} h_0 \\ h_1 \\ \vdots \\ h_M \end{bmatrix} = \begin{bmatrix} ? \\ ? \\ \vdots \\ ? \end{bmatrix}$$

for certain numbers “?” and determine all these “?”

- (b) Chapter 8 makes the point that  $u_t$  should be “sufficiently rich” (final part of § 8.1.3). Can you argue from the claim of Problem 3.10(c) of the lecture notes—page 26—that sufficiently rich ensures that the above equation has a unique solution  $h_0, \dots, h_M$ ?

problem:	1	2	3	4	5	6	7
points:	3+2+2	2+1+3+2+2+2	3	1	4	4	3+2

Exam grade is  $1 + 9p/p_{\max}$ .